

Solutions to tutorial exercises for stochastic processes

T1. \Rightarrow : Suppose \mathbb{X} is $\mathfrak{F} - \mathfrak{S}^T$ measurable. For any $t \in T$ we have by the definition of \mathfrak{S}^T that $\mathbb{X}^{-1}(\Pi_t^{-1}(S)) \in \mathfrak{F}$ for any $S \in \mathfrak{S}$, where Π_t denotes the projection on the t th coordinate. Finally $\mathbb{X}^{-1}(\Pi_t^{-1}(S)) = \{\omega : X_t(\omega) \in S\} = X_t^{-1}(S) \in \mathfrak{F}$.

\Leftarrow : Suppose all projections X_t are $\mathfrak{F} - \mathfrak{S}$ -measurable. Let $S \in \mathfrak{S}$ and $t \in T$. Then $\mathbb{X}^{-1}(\Pi_t^{-1}(S)) = X_t^{-1}(S) \in \mathfrak{F}$. So \mathbb{X} is measurable on the set $\{\Pi_t^{-1}(S) : t \in T, S \in \mathfrak{S}\}$. This set generates \mathfrak{S}^T , so \mathbb{X} is $\mathfrak{F} - \mathfrak{S}^T$ measurable.

T2. Since X_t is continuous and since \mathbb{Q} is dense in \mathbb{R} we have that

$$\sup_{t \in \mathbb{R}} X_t = \sup_{t \in \mathbb{Q}} X_t.$$

Let $a \in \mathbb{R}$. Then

$$\left\{ \sup_{t \in \mathbb{R}} X_t \leq a \right\} = \left\{ \sup_{t \in \mathbb{Q}} X_t \leq a \right\} = \bigcap_{t \in \mathbb{Q}} \{X_t \leq a\} \in \mathfrak{F}.$$

So $\sup_{t \in \mathbb{R}} X_t$ is measurable on the set $\{(-\infty, a] : a \in \mathbb{R}\}$, which generates \mathfrak{B} . So $\sup_{t \in \mathbb{R}} X_t$ is $\mathfrak{F} - \mathfrak{B}$ -measurable. Furthermore

$$\left\{ \sup_{t \in \mathbb{R}} X_t = \infty \right\} = \bigcap_{n=1}^{\infty} \bigcup_{t \in \mathbb{Q}} \{X_t \geq n\} \in \mathfrak{F},$$

and similarly

$$\left\{ \sup_{t \in \mathbb{R}} X_t = -\infty \right\} = \bigcap_{n=1}^{\infty} \bigcup_{t \in \mathbb{Q}} \{X_t \leq -n\} \in \mathfrak{F},$$

so that the events $\{\sup_{t \in \mathbb{R}} X_t = \infty\}$ and $\{\sup_{t \in \mathbb{R}} X_t = -\infty\}$ are measurable as well.

T3. We first show by induction that for some $s > 0$, N_s is Poisson distributed with parameter λs . Firstly $\mathbb{P}(N_s = 0) = e^{-\lambda s}$. Now suppose that $\mathbb{P}(N_s = k) = e^{-\lambda s} \frac{(\lambda s)^k}{k!}$ for all $s > 0$. Then by conditioning on τ_1 we find

$$\begin{aligned} \mathbb{P}(N_s = k+1) &= \int_0^s \lambda e^{-\lambda x} \mathbb{P}(N_{s-x} = k) dx = \int_0^s \lambda e^{-\lambda x} e^{-\lambda(s-x)} \frac{(\lambda(s-x))^k}{k!} dx \\ &= e^{-\lambda s} \frac{(\lambda s)^{k+1}}{(k+1)!}. \end{aligned}$$

So N_s is indeed Poisson distributed with parameter λs . Let $T_k := \sum_{i=1}^k \tau_i$ be the sequence of arrivals. We can write

$$N_s = \sum_{k=1}^{\infty} \mathbb{1}_{\{T_k \leq s\}},$$

and similarly

$$N_t - N_s = \sum_{k=N_s+1}^{\infty} \mathbb{1}_{\{T_k \leq t\}} = \sum_{k=1}^{\infty} \mathbb{1}_{\{T_{N_s+k} \leq t\}}.$$

We know that $T_{N_s+1} > s$. In fact by the memorylessness of the exponential distribution we have $T_{N_s+1} \sim s + \text{EXP}(\lambda)$. Similarly $T_{N_s+k} \sim s + T'_k$, where T'_k is an i.i.d. copy of T_k . Furthermore T_{N_s+k} is independent of N_s , since N_s is independent of $\tau_{N_s+1}, \tau_{N_s+2}, \dots$. Finally we have

$$\begin{aligned} \mathbb{P}(N_s = x, N_t - N_s = y) &= \mathbb{P}(N_s = x) \mathbb{P}(N_t - N_s = y \mid N_s = x) \\ &= \mathbb{P}(N_s = x) \mathbb{P}\left(\sum_{k=1}^{\infty} \mathbb{1}_{\{T_{N_s+k} \leq t\}} = y \mid N_s = x\right) \\ &= \mathbb{P}(N_s = x) \mathbb{P}\left(\sum_{k=1}^{\infty} \mathbb{1}_{\{s+T'_k \leq t\}} = y\right) \\ &= e^{-\lambda s} \frac{(\lambda s)^x}{x!} e^{-\lambda(t-s)} \frac{(\lambda(t-s))^y}{y!}. \end{aligned}$$

T4. The process $(M + N)_t$ is increasing and right-continuous, since M_t and N_t are increasing and right-continuous. Furthermore

$$(M + N)_t - (M + N)_s = M_t - M_s + N_t - N_s \sim \text{POI}((\lambda + \mu)(t - s)),$$

since $M_t - M_s$ and $N_t - N_s$ are independent and Poisson distributed with parameter $\lambda(t-s)$ and $\mu(t-s)$ respectively. It remains to show that $(M + N)_t$ has steps of size 1 almost surely. Construct the process M'_t by placing $X_i \sim \text{POI}(\lambda)$ points, x_1^i, \dots, x_k^i , uniformly at random in the interval $[i, i+1)$, so that $M'_t \stackrel{d}{=} M_t$. Similarly construct $N'_t \stackrel{d}{=} N_t$ by placing the points y_1^i, \dots, y_l^i in the interval $[i, i+1)$. Then $(M' + N')_t \stackrel{d}{=} (M + N)_t$. Suppose $(M' + N')_t$ has a jump of size 2. Then there exists an interval $[i, i+1)$ such that $x_v^i = y_w^i$ for some $v, w \in \mathbb{N}$. Now,

$$\begin{aligned} \mathbb{P}((M' + N')_t \text{ has jump of size 2}) &\leq \sum_{i=0}^{\infty} \sum_{k=0}^{\infty} \sum_{l=0}^{\infty} \sum_{v=1}^k \sum_{w=1}^l \mathbb{P}(X_i = k, Y_i = l, x_v^i = y_w^i) \\ &= \sum_{i=0}^{\infty} \sum_{k=0}^{\infty} \sum_{l=0}^{\infty} \sum_{v=1}^k \sum_{w=1}^l \mathbb{P}(X_i = k) \mathbb{P}(Y_i = l) \mathbb{P}(x_v^i = y_w^i) \\ &= 0, \end{aligned}$$

since $\mathbb{P}(x_v^i = y_w^i) = 0$. So $(M' + N')_t$ has steps of size 1 almost surely and thus $(M + N)_t$ as well.